# Lecture 6a: Dirichlet distribution; drawing from the posterior 

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## Likelihood for population probabilities of bins

- In an example in the last lecture, we observed that, conditional on the sample and using a flat prior, the probability distribution of the population probability of a single bin with $n$ observations and a sample size of $N$ is $\operatorname{Beta}(n+1, N-n+1)$
- This result generalizes: If we have a set of $k$ non-overlapping bins, their population probabiities given the sample are distributed as Dirichlet $\left(n_{1}, \ldots, n_{k}, n_{0}\right)$, where $n_{i}$ is the population count in bin $i$ for $i>0$ and $n_{0}=N-\sum_{1}^{k} n_{k}$.
- The Dirichlet pdf, with parameter vector $\vec{\alpha}$ is proportional to

$$
\prod_{i=1}^{k-1} p_{i}^{\alpha_{i}-1} \cdot\left(1-\sum_{1}^{k-1} p_{i}\right)^{\alpha_{k}} .
$$

Note that, though it has a parameter vector of length $k$, it is a pdf only for $k-1 p_{i}$ 's. Or, you can think of it as a pdf for a $k$-dimensional $\vec{p}$ on the unit simplex, the set of $\vec{p} \mathrm{~s}$ satisfying $\sum p_{i}=1$.

## More on the Dirichlet

- The conditional distribution of any collection of elements of $\vec{p}$, divided by their sum, conditional on the remaining elements of $\vec{p}$, is again Dirichlet, with parameter vector given by the corresponding elements of the original $\vec{\alpha}$.
- Since this conditional distribution does not depend on the other elements of $\vec{p}$, the normalized subvector is independent of the remaining elements of $\vec{p}$. Therefore the marginal distribution of the normalized subvector is the same Dirichlet.


## Using Monte Carlo draws from the posterior

- When the posterior density is complicated, or when we want to calculate the distribution or expectation of a function of the parameter that would require evaluating a nasty integral, it can be helpful to generate random draws from the posterior distribution.
- For example, suppose we have two bins $A$ and $B$ with modest and similar counts in the sample. A has more observations than $B$ in the sample, but we're not sure whether that's strong evidence that the underlying population probability of $A$ is bigger than that of $B$.
- The results above on the Dirichlet imply that the joint distribution of the two bin probabilities, $p_{A}$ and $p_{B}$, normalized to sum to one, is.
under a flat prior, proportional to $p_{A}^{n_{A}} p_{B}^{n_{B}}$, which is proportional to a $\operatorname{Beta}\left(n_{A}+1, n_{B}+1\right)$ density on $p_{A}$.
- To estimate the probability, given the data, that $p_{A}>p_{B}$, we can make, say, 1000 draws from a $\operatorname{Beta}\left(n_{A}+1, n_{B}+1\right)$ distribution (e.g. using rbeta() in R ) and count the proportion of the draws that have $p_{A}>.5$.
- Of course in this case we could also just look up the answer as qbeta (. $5, \mathrm{nA}+1, \mathrm{nB}+1$ ), but on the exercise for this week you'll check a probability that's more difficult to calculate analytically.

