

## INDEPENDENCE, TRANSFORMATIONS AND JACOBIANS, SIMULATION

### 1. REVIEW OF INDEPENDENCE

- If two random vectors  $X$  and  $Y$  have joint pdf  $p(x, y)$ , they are **independent** if and only if  $p(x, y) = q_X(x)q_Y(y)$ , where  $q_X$  and  $q_Y$  both integrate to one.
- In this case it is easy to verify that  $q_X$  and  $q_Y$  are the marginal pdf's of  $X$  and  $Y$  and also  $q_X(x) = q_{X|Y}(x|y)$ ,  $q_Y(y) = q_{Y|X}(y|x)$ , that is,  $q_X$  and  $q_Y$  are also the conditional pdf's of  $X|Y$  and  $Y|X$ .
- Obviously this means that the conditional distribution of  $\{Y|X\}$  does not depend on  $X$  and for any function  $f$  of  $Y$ ,  $E[f(Y)|X] = E[f(Y)]$ . (Of course also the same things with the  $Y, X$  roles reversed.)
- A more general definition:  $Y$  is independent of  $X$  if for every function  $g(Y)$  such that  $E[|g(Y)|] < \infty$ ,  $E[g(Y)|X] \equiv E[g(Y)]$ . It turns out that if this is true, the same is true with the roles of  $x$  and  $y$  reversed.
- Yet another definition:  $Y$  and  $X$  are independent if and only if we can write their joint cdf  $F_{X,Y}(x, y) = F_X(x)F_Y(y)$ .
- A collection  $\{X_1, \dots, X_n\}$  of random vectors is **mutually independent** if for every  $i$  and for every  $g$  with  $E[g(X_i)]$  defined and finite,  $E[g(X_i)|X_{-i}] = E[g(X_i)]$ . Here we're using the notation that  $X_{-i}$  means all the elements of the  $X$  vector except the one with index  $i$ . If they have a joint pdf, this is equivalent to

$$p(x_1, \dots, x_n) = \prod_{i=1}^n q_i(x_i).$$

- It is possible to have  $X_i$  independent of  $X_j$  for any  $i \neq j$  between 1 and  $n$ , yet to have the collection  $\{X_1, \dots, X_n\}$  not mutually independent. That is, pairwise independence does not imply mutual independence.

### 2. TRANSFORMATIONS AND JACOBIANS

Suppose we start with a random variable  $X$  with known pdf  $p(x)$ , but now want to find the pdf of  $Z = g(X)$ .

- $g$  had better be monotone, or else we have to break up its domain into pieces on which it is monotone, then sum up the results.
- Simply substituting  $g^{-1}(z) = x$  into  $g$  to obtain  $p(g^{-1}(z))$  produces the pdf of  $z$  only under highly restrictive conditions. These conditions occur fairly often in regression models, so it is all too common for people to forget that they do not hold generally.

- The correct formula: The pdf of  $Z = g(X)$  is given by

$$q(z) = p(g^{-1}(z)) \left| \frac{dx}{dz} \right| = p(g^{-1}(z)) \left| \frac{1}{g'(g^{-1}(z))} \right|.$$

- The formula is less messy if we can solve for  $h(z) = g^{-1}(z)$ . Then it is

$$q(z) = p(h(z)) |h'(z)|.$$

- A possibly helpful mnemonic device: always think of a pdf  $p(x)$  as  $p(x)dx$ . Then the Jacobian rule is

$$p(x)dx \rightarrow p(x(z)) \left| \frac{dx}{dz} \right| dz,$$

and the Jacobian looks like a natural correction for the fact that we are replacing  $dx$  with  $dz$ .

### 3. MULTIVARIATE GENERALIZATION

- In the univariate case, the  $|dx/dz|$  term is not usually called a “Jacobian”. That term comes from the multivariate case.
- With  $X$  a vector and  $X = h(Z)$  a vector-valued function with the dimensions of  $Z$  and  $X$  matching, the formula becomes

$$q(z) = p(h(z)) \text{abs} \left( \left| \frac{dx}{dz} \right| \right).$$

- The  $|dx/dz|$  term is what is properly called a Jacobian. It is defined as

$$\left| \frac{dx}{dz} \right| = \left| \left[ \frac{\partial h_i(z)}{\partial z_j} \right] \right|$$

### 4. EXAMPLES: SQUARING A NORMAL

- $X \sim N(0, 1)$ ,  $Z = \frac{1}{2}X^2$ .

$$p(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2}, \quad h(z) = \sqrt{2z}, \quad q(z) = \frac{1}{\sqrt{z\pi}} e^{-z}.$$

- What happened to the  $\frac{1}{2}$ ?
- This  $q$  is a special case of the general Gamma( $n, \alpha$ ), which has pdf

$$\text{Gamma}(z | n, \alpha) = \frac{1}{\Gamma(n)} \alpha^n z^{n-1} e^{-\alpha z}.$$

- The gamma function is defined for  $n > 0$ , satisfies  $\Gamma(n) = (n-1)!$  for integer  $n$  and  $n\Gamma(n) = \Gamma(n+1)$  for all  $n > 0$ .
- If  $Z \sim \text{Gamma}(n, \alpha)$ , then its mean  $E[Z]$  is  $n/\alpha$ .

- Notation for the gamma distribution is not well standardized. What we have called  $n$  is sometimes called  $p$  or  $\alpha$ . It is the “shape” parameter. What we have called  $\alpha$  is sometimes replaced by its inverse, which is then sometimes called  $\beta$  or  $\sigma$ .  $\alpha$  is the “inverse-scale” parameter, and its inverse is called the “scale”.

#### 5. EXAMPLES: SUM OF TWO INDEPENDENT RANDOM VARIABLES

- $Z = X + Y$ ,  $X$  and  $Y$  independent, pdf's  $p(x)$  and  $q(y)$ .
- To keep the Jacobian square, we have to think of the transformation as

$$(Z, X) = g(Y, X) \quad (Y, X) = h(Z, X).$$

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$$\frac{\partial h}{\partial z, x} = \left( \frac{\partial g}{\partial (y, x)} \right)^{-1} = \begin{bmatrix} 1 & 1 \\ 0 & 1 \end{bmatrix}^{-1}.$$

- So the Jacobian part is easy. But now we are left with the joint pdf  $p(x)q(z-x)$ . To get the marginal pdf  $r(z)$  of  $Z$  we have to integrate:

$$r(z) = \int_{-\infty}^{\infty} p(x)q(z-x) dx = p * q(z).$$

- This is called the **convolution** of  $p$  with  $q$ .
- These steps — Pad the transformation until it is square, apply Jacobian rule, integrate out parts of the transformed vector we are not interested in — are needed whenever we are deriving the pdf of a function of a vector of random variables that is of lower dimension than its arguments. In more complicated cases, these calculations can be burdensome.

#### 6. EXAMPLES: SUMS OF SQUARES OF NORMALS, SUMS OF GAMMAS

- $X$  and  $Y$  independent, both  $N(0, 1)$ .  $Z = X^2 + Y^2$ . We could do this directly, but note that we already know that  $U = \frac{1}{2}X^2$  and  $V = \frac{1}{2}Y^2$  are  $\text{Gamma}(\frac{1}{2}, 1)$ .
- So let's apply the convolution rule to the sum of two i.i.d. (independent, identically distributed) gammas,  $W=U+V$ :

$$q(w) = \int_0^w \frac{1}{\pi} u^{-\frac{1}{2}} (w-u)^{-\frac{1}{2}} e^{-w} du$$

- We can make the integral come out as a constant by replacing  $u$  with  $t = u/w$ . This leads to

$$q(w) = \int_0^1 \frac{1}{w\pi} t^{-\frac{1}{2}} (1-t)^{-\frac{1}{2}} e^{-w} w dt.$$

- Here, as often in dealing with pdf's, we don't need to determine the exact value of the integral now that we know it's a constant, call it  $\kappa$ , because the constant will be absorbed in the normalizing constant.

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$$q(w) = \kappa e^{-w} = \text{Gamma}(w | 1, 1).$$

- This same sort of trick gives the conclusion that in general the sum of a  $\text{Gamma}(n, \alpha)$  with a  $\text{Gamma}(m, \alpha)$  is a  $\text{Gamma}(n + m, \alpha)$ .
- In our original case now, we know that  $\frac{1}{2}Z \sim \text{Gamma}(1, 1)$ . This result can be applied recursively to get the conclusion that half the sum of  $n$  squared i.i.d.  $N(0, 1)$  variables is distributed as  $\text{Gamma}(\frac{n}{2}, 1)$ . Though the transformation from  $Z$  to  $\frac{1}{2}Z$  is trivial, the distribution of the sum of  $n$  squared i.i.d.  $N(0, 1)$  variates occurs so often that it is given a special name:  $\chi^2(n)$ . So obviously if  $X \sim \chi^2(n)$ ,  $\frac{1}{2}X \sim \text{Gamma}(\frac{n}{2}, 1)$ .

#### 7. EXAMPLES: WHERE A JACOBIAN CALCULATION MIGHT BE NASTY

- A common form of model is one that involves linear equations, written in matrix form as

$$Y\Gamma = XB.$$

- Often  $\Gamma$  is square and non-singular and we need to discuss how  $Y$  responds to  $X$ , which corresponds to finding  $\Pi$  in

$$Y = XB\Gamma^{-1} = X\Pi.$$

- Suppose we know the joint pdf's of the elements of  $\Gamma$  and  $B$ . What is the joint pdf of the elements of  $\Pi$ ? A terrible mess.

#### 8. SIMULATION: THE REFUGE OF THOSE WHO WANT TO AVOID CALCULUS

- Instead of taking derivatives of inverses of matrices, then integrating out redundant variables, we can simulate.
- Get the computer to generate a large sample of (pseudo-)random numbers with the joint distribution of the elements of  $\Gamma, B$ . For each draw  $i$ , form  $\Pi_i = B_i\Gamma_i^{-1}$ .
- If, for example, we want to know the expectation of  $\pi_{jk}$ , the  $j$ 'th row,  $k$ 'th column element of  $\Pi$ , we can just take the sample average of that element over our artificial sample  $\{\Pi_i\}$ , i.e. if there are  $N$  matrices in our sample, for  $N^{-1} \sum_i \pi_{jki}$ .
- To estimate the density function for  $\pi_{jk}$ , we can form a histogram or kernel density estimate from the artificial sample (`hist` or `density` in R).
- You do have to program the computer. It may take the computer a while, in a real-world application, to generate a big enough sample to give you accurate estimates. But you can spend the time productively elsewhere, and in some cases the computer will be faster than your calculus.